

# Comments on “Information Nonanticipative Rate Distortion Function And Its Applications”

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## Abstract

In [1, Theorem III.6] it is claimed that, for a one-sided random source  $x_1^\infty = x_1, x_2, \dots$ , the search for the non-anticipative (i.e., causal) rate distortion function can be restricted to reconstructions  $y_1^\infty$  which are jointly stationary with  $x_1^\infty$ . In this technical report we show that the proof of [1, Theorem III.6] is invalid because it relies on [1, Theorem III.5], the proof of which, as we also show, is flawed.

## I. INTRODUCTION

The manuscript [1] utilizes [2, Theorem 4] to prove the claim that, for one-sided sources  $x_1^\infty$ , the non-anticipative (i.e., causal) rate-distortion function can be realized by a reconstruction process  $y_1^\infty$  which is jointly stationary with  $x_1^\infty$ . To do so, it relies on [1, Theorem III.5].

In this note we argue that the proof of [1, Theorem III.5], and hence that of [1, Theorem III.6], are flawed. For that purpose, we will first recall the assumptions and definitions utilized in [2]. After that, we will present the definitions introduced in [1] and show, under the conditions stated there, the requirements needed by [2, Theorem 4] (the basis of [1, Theorem III.6]) of are not met.

## II. A BRIEF REVIEW OF [2]

Throughout [2], the search in the infimizations associated with various types of “nonanticipatory” (i.e., causal) rate-distortion functions is stated over sets of *joint* probability distributions between source and reconstruction (as opposed to the usual definitions, in which the search is over *conditional* distributions, see [3, Chapter 10], [4]). Since the distribution of the source is given, it is required that for every  $k_2 > k_1 \in \mathbb{Z}$ , all the joint distributions  $P_{x_{k_1}^{k_2}, y_{k_1}^{k_2}}$  to be considered yield  $x_{k_1}^{k_2}$  having the

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same (given) distribution of the source for the corresponding block, say  $P_{\mathcal{X}_{k_1}^{k_2}}$ . This requirement can be formalized as requiring that  $P_{\mathcal{X}_{k_1}^{k_2}, \mathcal{Y}_{k_1}^{k_2}} \in \mathcal{P}^{k_1, k_2}$ , for a set of admissible joint distributions  $\mathcal{P}^{k_1, k_2}$  defined as

$$\mathcal{P}^{k_1, k_2} \triangleq \left\{ P : P(E \times \mathcal{Y}_{k_1}^{k_2}) = P_{\mathcal{X}_{k_1}^{k_2}}(E), \quad \forall E \in \mathcal{B}(\mathcal{X}_{k_1}^{k_2}) \right\}, \quad k_1 \leq k_2 \in \mathbb{Z}, \quad (1)$$

where  $\mathcal{X}_{k_1}^{k_2}$  and  $\mathcal{Y}_{k_1}^{k_2}$  are, respectively, the alphabets to which  $x_{k_1}^{k_2}$  and  $y_{k_1}^{k_2}$  belong, and  $\mathcal{B}(\mathcal{X}_{k_1}^{k_2})$  is a  $\sigma$ -algebra over  $\mathcal{X}_{k_1}^{k_2}$ . In [2], this admissibility requirement is embedded in the definition of the sets of distributions which meet the distortion constraint, described next.

The fidelity criterion for every pair of integers<sup>1</sup>  $k_1 \leq k_2$  is expressed in [2] as requiring  $P_{\mathcal{X}_{k_1}^{k_2}, \mathcal{Y}_{k_1}^{k_2}}$  to belong to a non-empty set of distributions (hereafter referred to as *distortion-feasible set*)  $\mathcal{W}_D^{k_1, k_2}$ , a condition written as  $(x_{k_1}^{k_2}, y_{k_1}^{k_2}) \in (\mathcal{W}_D^{k_1, k_2})$ . In this definition, the number  $D \geq 0$  represents an admissible distortion level. Notice that such general formulation of a fidelity criteria does not need a distortion function and does not necessarily involve an expectation.

As mentioned above, the admissibility requirement  $P_{\mathcal{X}_{k_1}^{k_2}} \in \mathcal{P}^{k_1, k_2}$  is expressed in the distortion-feasible sets in [2, eqn. (2.1)]. The latter equation can be written as

$$\mathcal{W}_D^{k_1, k_2} \subset \mathcal{P}^{k_1, k_2}. \quad (2)$$

In [2, eqs. (2.4) and (2.5)], the distortion-feasible sets are assumed to satisfy the ‘‘concatenation’’ condition

$$(x_{k_1}^{k_2}, y_{k_1}^{k_2}) \in (\mathcal{W}_D^{k_1, k_2}) \wedge (x_{k_2+1}^{k_3}, y_{k_2+1}^{k_3}) \in (\mathcal{W}_D^{k_2+1, k_3}) \implies (x_{k_1}^{k_3}, y_{k_1}^{k_3}) \in (\mathcal{W}_D^{k_1, k_3}). \quad (3)$$

With this, [2, eqn. (2.9)] defined the ‘‘nonanticipatory epsilon entropy’’ of the set of distributions<sup>2</sup>  $\mathcal{W}_D^{k_1, k_2}$  as

$$H^0(\mathcal{W}_D^{k_1, k_2}) \triangleq \inf I(x_{k_1}^{k_2}; y_{k_1}^{k_2}), \quad (4)$$

where the infimum is taken over all pairs of random sequences  $(x_{k_1}^{k_2}, y_{k_1}^{k_2}) \in (\mathcal{W}_D^{k_1, k_2})$  such that the causality Markov chains

$$x_{k+1}^{k_2} \longleftrightarrow x_{k_1}^k \longleftrightarrow y_{k_1}^k, \quad k_1 \leq k \leq k_2 \quad (5)$$

are satisfied. Then [2, eq. (2.13)] defines the ‘‘nonanticipatory message generation rate’’ as

$$\overline{H}_D^0 \triangleq \lim_{k_2 - k_1 \rightarrow \infty} \frac{1}{k_2 - k_1} H^0(\mathcal{W}_D^{k_1, k_2}) \quad (6)$$

<sup>1</sup>The analysis in [2] considered both discrete- and continuous-time processes, but here we only refer to the discrete-time scenario.

<sup>2</sup>The actual term employed in [2] is ‘‘nonanticipatory epsilon entropy of the message  $(\mathcal{W}_D^{k_1, k_2})$ ’’ where the term ‘‘message’’ refers to the random ensembles in  $(\mathcal{W}_D^{k_1, k_2})$ .

(when the limit exists).

An alternative “nonanticipatory message generation rate” is also considered in [2] by defining the set of distortion-admissible process distributions  $\mathcal{W}_D$  as follows:

**Definition 1.** *The set  $(\mathcal{W}_D)$  consists of all two-sided random process pairs  $(x_{-\infty}^{\infty}, y_{-\infty}^{\infty}) \in (\mathcal{W}_D)$  for which there exist integers  $\dots < k_{-1} < k_0 < k_1 < \dots$  such that  $\lim_{i \rightarrow \pm\infty} k_i = \pm\infty$  and*

$$(x_{k_i}^{k_{i+1}-1}, y_{k_i}^{k_{i+1}-1}) \in (\mathcal{W}_D^{k_i, k_{i+1}-1}), \quad \forall i \in \mathbb{Z}. \quad (7)$$

▲

With this, [2, eq. (2.14)] defines

$$\overrightarrow{H}_D^0 \triangleq \inf \lim_{k_2 - k_1 \rightarrow \infty} \frac{1}{k_2 - k_1} I(x_{k_1}^{k_2}; y_{k_1}^{k_2}) \quad (8)$$

(when the limit exists), where the infimum is taken over all pairs of processes  $(x_{-\infty}^{\infty}, y_{-\infty}^{\infty}) \in (\mathcal{W}_D)$  satisfying the causality Markov chains

$$x_{k+1}^{\infty} \longleftrightarrow x_{-\infty}^k \longleftrightarrow y_{-\infty}^k, \quad \forall k \in \mathbb{Z}. \quad (9)$$

### III. THE PROBLEMS WITH [1]

The proof of [1, Theorem III.6] relies on the claim stated in [1, Theorem III.5], namely, that an equality similar to

$$\overrightarrow{H}_D^0 = \overline{H}_D^0 \quad (10)$$

holds.

We demonstrate that the proof of [1, Theorem III.5] is not valid (and hence that of [1, Theorem III.6] is flawed). We do this by showing next that [1, Theorem III.5] has two problems, namely: a) one of the causal IRDFs considered in it does not coincide with  $\overrightarrow{H}_D^0$ , and b) the proof of [1, Theorem III.5] is invalid.

#### A. The First Problem

The already mentioned first problem of [1, Theorem III.5] as a basis for [1, Theorem III.6] follows from the fact that [1] defines its alternative causal IRDF function  $\overrightarrow{R}^{na}(D)$  as ([1, II.9])

$$\overrightarrow{R}^{na}(D) \triangleq \inf_{P_{y_1^{\infty} | x_1^{\infty}} \in \overrightarrow{\mathcal{Q}}_{1, \infty}(D)} \lim_{n \rightarrow \infty} \frac{1}{n} I(x_1^n \rightarrow y_1^n), \quad (11)$$

where (as defined in the text just below equation (II.6) in [1])  $\overrightarrow{\mathcal{Q}}_{1, \infty}(D)$  is the set of conditional distributions of  $y_1^{\infty}$  given  $x_1^{\infty}$  such that  $(x_1^{\infty}, y_1^{\infty})$  satisfies the causality Markov chains

$$x_{k+1}^{\infty} \longleftrightarrow x_1^k \longleftrightarrow y_1^k, \quad k = 1, 2, \dots \quad (12)$$

and the asymptotic distortion constraint

$$\lim_{n \rightarrow \infty} \frac{1}{n} \mathbb{E}[d_{1,n}(x_1^n, y_1^n)] \leq D. \quad (13)$$

Next, [1] states in its equation (III.2) that [2] defined

$$\vec{R}^\varepsilon(D) \triangleq \inf_{P_{y_1^\infty | x_1^\infty} \in \vec{\mathcal{Q}}_{1,\infty}(D)} \lim_{n \rightarrow \infty} \frac{1}{n} I(x_1^n; y_1^n). \quad (14)$$

Thanks to (12), it readily follows that  $\vec{R}^\varepsilon(D) = \vec{R}^{na}(D)$  (although this equality is not explicitly stated in [1]).

Since the only causal IRDF defined in [2] as an inf lim is  $\vec{H}_D^0$ , one must conclude that [1] regards  $\vec{R}^\varepsilon(D)$  as equivalent to  $\vec{H}_D^0$ . However, in view of Definition 1 and (8), such equivalence is not valid (since the distortion feasible sets of Definition 1 are not compatible with the distortion constraint (13)). Therefore, when in [1, Theorem III.5] it is stated that  $R^{na}(D) = \vec{R}^{na}(D)$  (and hence  $\vec{H}_D^0 = \vec{R}^\varepsilon(D)$ ), it does not mean that  $\vec{H}_D^0$  equals  $\vec{H}_D^0$ . As a consequence, one of the necessary conditions for [2, Theorem 4] is not shown to hold.

### B. The Second Problem

The second issue with [1, Theorem III.5] is the validity of its proof. To begin with, the only argument used in it is that the source is stationary and [2, Theorem 2]. However, the latter theorem only says that  $\vec{H}_D^0 \leq \vec{H}_D^0$ , and thus the proof of [1, Theorem III.5] presented there is flawed.

Although not referred to in that proof, the reverse inequality claimed in [1, Lemma III.4] would be all that is required to show that  $\vec{R}^\varepsilon(D) = \vec{H}_D^0$ . However, the proof of [1, Lemma III.4], reproduced below, is clearly invalid. It starts by noting that, by definition,

$$R_{1,n}^{na}(D) \leq I(x_1^n, y_1^n), \quad \forall (x_1^n, y_1^n) \in (\vec{\mathcal{Q}}_{1,n}(D)). \quad (15)$$

Then it proceeds by saying that ‘‘taking the limit on both sides we obtain’’

$$\lim_{n \rightarrow \infty} \frac{1}{n} R_{1,n}^{na}(D) \leq \lim_{n \rightarrow \infty} \frac{1}{n} I(x_1^n, y_1^n), \quad \forall (x_1^\infty, y_1^\infty) \in (\vec{\mathcal{Q}}_{1,\infty}(D)) \quad (16)$$

and then that the claim follows by taking the infimum over  $\vec{\mathcal{Q}}_{1,\infty}(D)$ . The problem with this reasoning is that (16) does not follow from (15). A rigorous reasoning reveals that when taking the limit as  $n \rightarrow \infty$ , (15) translates to

$$\lim_{n \rightarrow \infty} R_{1,n}^{na}(D) \leq \lim_{n \rightarrow \infty} \frac{1}{n} I(\overset{(n)}{x}_1; \overset{(n)}{y}_1), \quad \forall \{\overset{(n)}{x}_1, \overset{(n)}{y}_1\}_{n \in \mathbb{N}} \text{ such that } P_{\overset{(n)}{y}_1 | \overset{(n)}{x}_1} \in \vec{\mathcal{Q}}_{1,n}(D) \quad (17)$$

Thus, one cannot choose to infimize the RHS of this inequality over  $\vec{\mathcal{Q}}_{1,\infty}(D)$  and expect the inequality to hold, since one can easily find a pair of processes  $(x_1^\infty, y_1^\infty)$  whose conditional distribution  $P_{y_1^\infty | x_1^\infty}$  belongs to  $\vec{\mathcal{Q}}_{1,\infty}(D)$  and yet  $P_{y_1^n | x_1^n} \notin \vec{\mathcal{Q}}_{1,n}(D)$  (because the normalized expectations on the LHS of (13) are allowed to reach the limit  $D$  from above).

In order to arrive to (16), one should first show that

$$R_{1,n}^{na}(D) \leq \frac{1}{n} I(x_1^n, y_1^n), \quad \forall P_{y_1^\infty | x_1^\infty} \in \vec{\mathcal{Q}}_{1,\infty}(D). \quad (18)$$

Unfortunately, the latter is not true since, as already mentioned,  $\vec{\mathcal{Q}}_{1,\infty}(D)$  allows pairs of random processes  $(x_1^\infty, y_1^\infty)$  such that  $\mathbb{E}[d_{1,n}(x_1^n, y_1^n)] > D$ , for all  $n \in \mathbb{N}$  (thus reaching the limit distortion (13) from above), and thus such that  $P_{y_1^n | x_1^n} \notin \vec{\mathcal{Q}}_{1,n}(D)$ , for all  $n \in \mathbb{N}$ . Therefore, (18) does not hold. Indeed, the latter reasoning reveals that

$$R_{1,n}^{na}(D) \geq \inf_{(x_1^\infty, y_1^\infty) \in (\vec{\mathcal{Q}}_{1,\infty}(D))} \frac{1}{n} I(x_1^n, y_1^n), \quad n \in \mathbb{N}, \quad (19)$$

leading to an inequality in the same direction as the one provided by [2, Theorem 2], i.e., that  $R^{na}(D) \geq \vec{R}^\varepsilon(D)$ .

## REFERENCES

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